Adopting spatial copula framework

in regional frequency analysis

STAPPY 200

Statistical Hydrology Workshop



M. Durocher(1), F. Chebana(1) and T.B.M.J. Ouarda(2) Institut National de Recherche Scientifique, 490 rue de la Couronne, Québec, Canada (martin.durocher@ete.inrs.ca, fateh.chebana@ete.inrs.ca) (2) Masdar Institue of Science and Technology P.O. Box 54224, Abu Dhabi, UAE (touarda@masdar.ac.ae)



Université d'avant-garde

1. Introduction

Context

- RFA is used for predicting flood quantiles at ungauged locations. The rational is that two sites having similar hydrological properties should behave similarly. In that case, valuable information can be transfered between sites.
- Hydrological dissimilarity between a gauged site and an ungauged location cannot be calculated. Instead, a physiographical space in which an **associated metric** between site characteristics must be used.
- In physiographical spaces, flood quantiles can be predicted by interpolation methods, such kriging.

Problematic

Usually, flood quantiles share a log-log relationship with site characteristics, which creates bias and suboptimal prediction with traditional kriging techniques.

2. Case study

Hydrological data

Specific flood quantiles of 100-year return period

At-site analysis

Southern Quebec,

Canada

Number of sites :151 Natural flow regime



- Traditional kriging does not account for non-constant variance.
- The problem associated with traditional kriging technique can be resolved by considering Spatial Copula, an extension of traditional geostatistical framework where spatial dependance is characterized by a copula.

2. Spatial copula

A multivarite distribution *G* can be expressed as

 $G(\mathbf{x}) = C\left[(F_1(x_1), \ldots, F_n(x_n))\right]$

where $\{F_i\}_{i=1}^n$ are margins for $\mathbf{x}' = (x_1, \ldots, x_n)$ and C is a copula

For spatial analysis, the copula has the same dimension as the number of site *n* and **the strengh of the dependance** must be associated with a **distance** *h*.

Spatial copula must allows for strong dependance

$$C_h \rightarrow M^n$$
 when $h \rightarrow 0$

and perfect independence

$$C_h \to \Pi^n$$
 when $h \to \infty$

▶ In spatial copula framework, the margins of the distribution G are treated separately from its dependence. Hence, there is 2 set of parameters : the marginal part η and the copula part θ . Joint estimation of both part can be performed by optimizing a pairwise likelihood

$$L(\mathbf{z} \mid \eta, \theta) = \prod_{i < j} f(z_i, z_j \mid \eta, \theta)$$

where $\mathbf{z}' = (z_1, \ldots, z_n)$ are spatial observation and *f* is the bivariate density of two sites *i* and *j*.

For known parameters $(\hat{\eta}, \hat{\theta})$, the plug-in predictive distribution (PPD) at ungauged location is the product of the marginal density and the conditional copula [2] :

$$p(z \mid \mathbf{z}, \theta) = f_{\hat{\eta}}(z) \times c_{\hat{\theta}} \left[F_{\hat{\eta}}^{-1}(z) \mid \mathbf{w} \right]$$

where $\mathbf{w}' = (w_1, \dots, w_n)$ and $w_i = F_{\hat{\eta}}^{-1}(z_i)$

Predictors can be calculated from the mean or the median of the PPD. For instance, the median is the quantity $F_{\hat{n}}^{-1}(w^*)$ for which

$$1/2 = \int_0^{w^*} c_{\hat{\theta}}(u \mid \mathbf{w}) du$$

Record length: >15 years

Fig 1 : Map of the stations

5. Results

The performance of the prediction obtained by spatial copula is assessed by Leave-one-out cross-validation. In turn each gauged station is considered as ungauged and a predicted value is obtained as the median of the PPD.

The analysis of the residuals shows the presence of large relative discrepancies (Fig. 5-Left). Note the presence of problematic stations previously identified for this database [3].

The absolute residuals at logarithm scale (Fig. 5-Right) show a decreasing variance that is coherent with the trend $\sigma(S_i)$



3. Physiographical space

There is no general agreement on what hydrological similarity between sites should be. Here the **dissimilarity** between two sites *i* and *j* is defined as the **hydrological distance**

 $h_{i,j} = d(Z_i, Z_j)$

between the vectors of flood quantiles $Z_i = (Z_{i,1}, \ldots, Z_{i,r})$ with return periods 1, ..., *r*

Let a physiographical space be a subspace of **coordinates**

 $S_i = AX_i$

where X_i are site characteristics and A is a matrix that projects X_i on the physiographical space.

A metric is associated with a dissimilarity measure if

> Small $d(S_i, S_j)$ \rightarrow Small $d(Z_i, Z_j)$

Prediction - 0.8 - 0.6 - 0.4 - 0.2 -2 2 3 Std. dev. - 0.20 - 0.15 - 0.10 - 0.05

4. Model

Marginal part(η)

Regional distribution of the flood quantiles is log-normal.

 $log(Z_i) \rightarrow N\left[\mu(S_i), \sigma^2(S_i)\right]$

for the strong correlation between

the first canonical coordinates $S_{i,1}$

 $\mu(S_i) = \beta_{\mu,0} + \beta_{\mu,1} S_{i,1}$

 $\sigma(\mathbf{S}_i) = \beta_{\sigma,0} + \beta_{\sigma,1} \mathbf{S}_{i,1}$

and the flood quantiles:



Fig. 3 : Normalized QQ-plot

Copula part (θ)

The spatial dependance is characterized by a Gaussian copula with pairwise correlation

$$ho(S_i, S_j \mid \lambda, \tau) = (1 - \tau) \exp\left[-3 rac{d(S_i, S_j)}{\lambda}
ight]$$

where $\lambda > 0$ (practical range) controls the correlation as $d(S_i, S_i) \rightarrow \infty$ and τ is a local measurement error (nugget effect)

-3.0 -2.5 -2.0 -1.5 -1.0 -0.5 -3.0 -2.5 -2.0 -1.5 -1.0 -0.5

Prediction (log)

Prediction (log)

Fig 5 : Residuals

The prediction power of the spatial copula approach (SCop) is compared to other methods on the same database:

- Multiple regression with CCA-delineation (CCA)
- Residual drift kriging (Krig)
- Generalized additive model (GAM)
- Artificial neural networks (ANN)



Fig 6 : Performance criteria

In comparison with traditional kriging (Krig), the results of spatial copula (Scop) is associated with an important reduction of the relative bias.

Fig 2 : Predictions in physio.

space

A way to build $A = (a_1, a_m)$ is by using canonical correlation analysis, for which canonical pairs

 $\mathbf{s}_k = \mathbf{a}_k X$ and $\mathbf{v}_k = \mathbf{b}_k Z$

sequentially optimize $corr(\mathbf{s}_k, \mathbf{v}_k)$ for k = 1, ..., m

The Goodness-of-fit test [1] based on binned pairwise observations validates the choice of a Gaussian copula. For each bins, P-values are larger than 20% are found.



Overall, Scop and ANN have the **best** rel.RMSE from the methods considered here.

7. Conclusion

The spatial copula framework offers a full probabilistic model that can account for non-constant variance.

The spatial copula framework appears more appropriate in presence of problematic stations in comparison with traditional kriging.

The spatial copula framework has competitive performance with the best method. In particular, it improves over traditional kriging.

The important relative bias observed with the traditional kriging approach is reduced greatly with the spatial copula approach.

References

- [1] Bárdossy, A. (2006) Copula-Based Geostatistical Models for Groundwater Quality Parameters. Water ressour res, 42(11).
- [2] Bárdossy, A., and Li, J. (2008) Geostatistical Interpolation Using Copulas. Water ressour res, 44(7).
- [3] Chokmani K, Ouarda TBMJ (2004) Physiographical space-based kriging for regional flood frequency estimation at ungauged sites. Water ressour res. 40(12).

Acknowledgments

Financial support was provided by the Natural Sciences and Engineering Research Council (NSERC) of Canada.